A calculator for mortality following emergency general surgery based on the American College of Surgeons National Surgical Quality Improvement Program database

Ivy N. Haskins, MD, Patrick J. Maluso, MD, Mary E. Schroeder, MD, Richard L. Amdur, PhD, Khashayar Vaziri, MD, Samir Agarwal, MD, and Babak Sarani, MD, Washington, District of Columbia

AAST Continuing Medical Education Article

Accreditation Statement

This activity has been planned and implemented in accordance with the Essential Areas and Policies of the Accreditation Council for Continuing Medical Education through the joint providership of the American College of Surgeons and the American Association for the Surgery of Trauma. The American College Surgeons is accredited by the ACCME to provide continuing medical education for physicians.

AMA PRA Category 1 CreditsTM

The American College of Surgeons designates this journal-based CME activity for a maximum of 1 AMA PRA Category 1 CreditTM. Physicians should claim only the credit commensurate with the extent of their participation in the activity.

Of the AMA PRA Category 1 Credit $^{\rm TM}$ listed above, a maximum of 1 credit meets the requirements for self-assessment.

Credits can only be claimed online



AMERICAN COLLEGE OF SURGEONS

Inspiring Quality: Highest Standards, Better Outcomes

100+years

Objectives

After reading the featured articles published in the Journal of Trauma and Acute Care Surgery, participants should be able to demonstrate increased understanding of the material specific to the article. Objectives for each article are featured at the beginning of each article and online. Test questions are at the end of the article, with a critique and specific location in the article referencing the question topic.

Claiming Credit

To claim credit, please visit the AAST website at http://www.aast.org/ and click on the "e-Learning/MOC" tab. You must read the article, successfully complete the post-test and evaluation. Your CME certificate will be available immediately upon receiving a passing score of 75% or higher on the post-test. Post-tests receiving a score of below 75% will require a retake of the test to receive credit.

Disclosure Information

In accordance with the ACCME Accreditation Criteria, the American College of Surgeons, as the accredited provider of this journal activity, must ensure that anyone in a position to control the content of *J Trauma Acute Care Surg* articles selected for CME credit has disclosed all relevant financial relationships with any commercial interest. Disclosure forms are completed by the editorial staff, associate editors, reviewers, and all authors. The ACCME defines a 'commercial interest' as "any entity producing, marketing, re-selling, or distributing health care goods or services consumed by, or used on, patients." "Relevant" financial relationships are those (in any amount) that may create a conflict of interest and occur within the 12'months preceding and during the time that the individual is engaged in writing the article. All reported conflicts are thoroughly managed in order to ensure any potential bias within the content is eliminated. However, if you'perceive a bias within the article, please report the circumstances on the evaluation form.

Please note we have advised the authors that it is their responsibility to disclose within the article if they are describing the use of a device, product, or drug that is not FDA approved or the off-label use of an approved device, product, or drug or unapproved usage.

Disclosures of Significant Relationships with Relevant Commercial Companies/Organizations by the Editorial Staff

Ernest E. Moore, Editor: PI, research support and shared U.S. patents Haemonetics; PI, research support, TEM Systems, Inc. Ronald V. Maier, Associate editor: consultant, consulting fee, LFB Biotechnologies. Associate editors: David Hoyt and Steven Shackford have nothing to disclose. Editorial staff: Jennifer Crebs, Jo Fields, and Angela Sauaia have nothing to disclose.

Author Disclosures

The authors have nothing to disclose.

Reviewer Disclosures

The reviewers have nothing to disclose.

Cost

For AAST members and *Journal of Trauma and Acute Care Surgery* subscribers there is no charge to participate in this activity. For those who are not a member or subscriber, the cost for each credit is \$25.

System Requirements

The system requirements are as follows: Adobe® Reader 7.0 or above installed; Internet Explorer® 7 and above; Firefox® 3.0 and above, Chrome® 8.0 and above, or SafariTM 4.0 and above.

Questions

If you have any questions, please contact AAST at 800-789-4006. Paper test and evaluations will not be accepted.

Submitted: November 4, 2016, Revised: February 15, 2017, Accepted: March 13, 2017, Published online: March 21, 2017.

From the Center for Trauma and Critical Care, Department of Surgery (I.N.H., P.J.M., M.E.S., R.L.A., K.V., B.S.), The George Washington University, Washington, District of Columbia; and Division of Colon and Rectal Surgery, Department of Surgery (S.A.), The George Washington University, Washington, District of Columbia.

Presented at the 102nd American College of Surgeons Clinical Congress, October 16th to 20th, 2016, in Washington, DC.

Address for reprints: Ivy N. Haskins, MD, 2150 Pennsylvania Ave, NW, Suite 6B, Washington, DC 20037; email: ihaskins@gwu.edu.

DOI: 10.1097/TA.0000000000001451

J Trauma Acute Care Surg Volume 82, Number 6 BACKGROUND: The complex nature of current morbidity and mortality predictor models do not lend themselves to clinical application at the bed-

side of patients undergoing emergency general surgery (EGS). Our aim was to develop a simplified risk calculator for prediction of

early postoperative mortality after EGS.

METHODS: EGS cases other than appendectomy and cholecystectomy were identified within the American College of Surgeons National Sur-

gery Quality Improvement Program database from 2005 to 2014. Seventy-five percent of the cases were selected at random for model development, whereas 25% of the cases were used for model testing. Stepwise logistic regression was performed for creation of a 30-day mortality risk calculator. Model accuracy and reproducibility was investigated using the concordance index (c sta-

tistic) and Pearson correlations.

RESULTS: A total of 79,835 patients met inclusion criteria. Overall, 30-day mortality was 12.6%. A simplified risk model formula was derived

from five readily available preoperative variables as follows: 0.034*age + 0.8*nonindependent status + 0.88*sepsis + 1.1 (if bun ≥ 29) or 0.57 (if bun ≥ 18 and < 29) + 1.16 (if albumin < 2.7), or 0.61 (if albumin ≥ 2.7 and < 3.4). The risk of 30-day mortality was stratified into deciles. The risk of 30-day mortality ranged from 2% for patients in the lowest risk level to 31%

for patients in the highest risk level. The c statistic was 0.83 in both the derivation and testing samples.

CONCLUSION: Five readily available preoperative variables can be used to predict the 30-day mortality risk for patients undergoing EGS. Further

studies are needed to validate this risk calculator and to determine its bedside applicability. (J Trauma Acute Care Surg. 2017;82:

1094–1099. Copyright $\ensuremath{\mathbb{C}}$ 2017 Wolters Kluwer Health, Inc. All rights reserved.)

LEVEL OF EVIDENCE: Prognostic/epidemiological study, level III.

KEY WORDS: Emergency; mortality; risk; surgery.

mergency general surgery (EGS) is an umbrella term that represents a heterogeneous collection of illnesses and patients connected only by their emergent nature. The diversity of EGS patients creates a clinical challenge due to the wide distribution of postoperative morbidity and mortality. Although many models exist to help predict patient morbidity and mortality, recent studies demonstrate that the estimates provided by these elective models may not accurately reflect the risk for patients undergoing EGS. Furthermore, many of these models require a large number of patient and operative variables which limits their application at the bedside of an EGS patient.

The volume of EGS patients in the United States is on the rise while access to such care continues to decline. Inherently, this has led to a shortage of medical resources with a lack of knowledge on how best to distribute them.³ In a study published by Gale et al.³ in 2014, the incidence of hospital admissions in 2010 for EGS was 1,290 admissions per 100,000 people, exceeding that of many common public health diseases, including coronary artery disease, all types of cancer, and new-onset diabetes mellitus.

Early mortality after EGS is estimated to be as high as 13% to 18%, a risk which is at least eight times higher than that associated with elective general surgery. 1,4,5 As the burden of EGS continues to grow, the need to accurately and efficiently risk stratify this heterogeneous patient population will be crucial to setting patient, family, and physician expectations. The purpose of our study was to develop a 30-day mortality risk calculator for patients undergoing EGS using the American College of Surgeons National Surgical Quality Improvement Program (ACS-NSQIP) database.

MATERIALS AND METHODS

All patients undergoing EGS from the years 2005 through 2014 were identified within the NSQIP database by Current Procedural Terminology codes. Specifically, all patients who underwent EGS for perforated peptic ulcer disease, small- or large-bowel disease requiring bowel resection, incarcerated or strangulated umbilical, inguinal, femoral, or ventral hernia

repair, and those patients who underwent debridement of necrotizing fasciitis were identified with the following Current Procedural Terminology codes: 43840, 44602, 49905, 44120, 44125, 44140, 44141, 44143, 44144, 44150, 44151, 44155, 44156, 44160, 49561, 49566, 49568, 49507, 49521, 49553, 49557, 49587, 11004, 11043, and 11046.

A procedure-specific calculator was not generated as we wanted to limit the complexity of the mortality risk model. Patients who underwent appendectomy and cholecystectomy were intentionally excluded from this analysis because these cases are often performed on an outpatient basis despite their urgent nature. Finally, encoded within the NSQIP database is an "emergency" variable which was required for case inclusion to ensure that nonemergent cases were excluded from our analysis.

Seventy-five percent of the cases identified within the NSQIP database were selected at random to be included in a derivation sample, whereas the remaining 25% of cases were used for model testing. To strengthen the accuracy of this model, only preoperative, objective variables were eligible for model inclusion. Laboratory values were coded into quartiles with "missing" being a fifth possible code.

Within the model derivation group, the association of preoperative patient demographics, comorbidities, and laboratory values with 30-day mortality was first investigated using univariate analysis. Student's t test was used for continuous variables, whereas χ^2 or Fisher's exact test was used for categorical variables. Variables that had a significant association with 30-day mortality at a p value less than 0.10 in univariate analysis were then included in a multivariate model. Multivariate logistic regression was used to identify the variables that had an independent association with 30-day mortality while controlling for relevant covariates. A backward-elimination approach was used in which predictors with p value greater than 0.10 were dropped at each step to allow for a final model.

All variables that had a significant association with 30-day mortality after multivariate logistic regression were used to generate an initial mortality risk calculator. This risk calculator was then simplified by eliminating those variables that had a weaker association with 30-day mortality using the

concordance index (c statistic) as a measure of model strength. The concordance index is a reflection of the probability that, for each pair of randomly chosen patients, the patient with a higher risk score for mortality is the patient that is less likely to survive. A value for the c-index greater than 0.8 indicates good discrimination. This means that the patient with the higher risk of mortality is the one who is accurately predicted to die 80% of the time. Variables were ranked based on the parameter estimate divided by the standard error, where a larger score indicates a more important predictor. Less important predictors were removed with the c statistic reexamined at each step. We kept variables in the final model that allowed the c statistic to drop by no more than 0.03 points while remaining greater than 0.8.

The simplified mortality risk model was used to calculate a risk score for each subject within both the model derivation and model testing groups using the same regression equation. Following the standard practice for assessing model testing, the mortality risk scores were stratified into deciles. The calibration of the model was assessed in both the derivation and testing data sets by examining the predicted versus observed incidence of 30-day mortality in each decile using the coefficient of determination (R^2) . In a perfectly calibrated model, the plot of predicted versus observed incidence of 30-day mortality across deciles should form a regression line with slope = 1.0 and y-intercept = 0. The reproducibility of the risk model was investigated by examining Pearson correlations between model parameters in the derivation and testing data sets. Model discrimination was examined using c statistics in both data sets.

SAS version 9.3 (Cary, NC) was used for all data analyses with p values less than 0.05 considered significant.

RESULTS

A total of 79,835 patients underwent EGS from the years 2005 through 2014; 59,716 of these cases were included in the derivation sample, whereas 20,119 of these cases were included in the testing sample. With respect to preoperative patient demographics, the average age of patients in the derivation group was 63.1 years, whereas the average of patients in the testing group was 63.3 years. Most patients in both groups were women (52.2% in the derivation group vs 52.6% in the validation group).

The association of preoperative patient variables with 30-day mortality in the derivation group is outlined in Table 1. All variables except sex were significantly associated with 30-day mortality. Although the definition of preoperative sepsis is available within the ACS-NSQIP Participant User File, we believe that it is important to define. For this study, preoperative sepsis was defined as any patient who met criteria for a systemic inflammatory response plus either a positive blood culture or infection site culture or preoperative presence of an infection as the indication for surgery, that is, systemic inflammatory response plus bowel strangulation. ¹⁰

Next, the association of preoperative patient laboratory values with 30-day mortality was investigated in univariate analysis. As mentioned previously, to perform this analysis, patients were divided into quartiles, with quartile 1 corresponding to the lowest laboratory values and quartile 4 corresponding to the highest laboratory values present in the dataset. Patients with a higher white blood cell count (p < 0.0001), a lower hematocrit level (p < 0.0001), a lower platelet count (p < 0.001), a lower sodium level (p < 0.001), a higher blood urea nitrogen level (BUN; p < 0.001), a higher creatinine level (p < 0.0001), a lower serum

TABLE 1. Association of Preo	perative Patient Variables With 30-D	av Mortality, Derivation Group

Variable	All Patients (N = 59,716)	Patients Dead at 30 d (N = 7,503)	Patients Alive at 30 d (N = 52,213)	p
Age (mean, SD), y	63 ± 17	72 ± 13	62 ± 17	< 0.0001
Female sex (N, %)	31,155 (52%)	3,983 (53%)	27,172 (52%)	0.15
Body mass index (mean, SD), kg/m ²	29 ± 9	28 ± 10	29 ± 9	< 0.0001
Nonindependent functional status (N, %)	11,247 (19%)	3,429 (31%)	4,074 (8%)	< 0.0001
Current smoker (N, %)	14,108 (24%)	1,638 (22%)	12,470 (24%)	< 0.0001
Hypertension (N, %)	32,467 (54%)	5,240 (70%)	27,227 (52%)	< 0.0001
CHF (N, %)	1,975 (3%)	746 (10%)	1,229 (2%)	< 0.0001
Dyspnea on exertion (N, %)	7,639 (13%)	2,125 (28%)	5,514 (11%)	< 0.0001
COPD (N, %)	6,316 (11%)	1,627 (22%)	4,689 (9%)	< 0.0001
Ascites (N, %)	3,694 (6%)	1,035 (14%)	2,659 (5%)	< 0.0001
ESRD (N, %)	2,254 (4%)	799 (36%)	1,455 (3%)	< 0.0001
Acute kidney injury (N, %)	2,324 (4%)	838 (11%)	1,486 (3%)	< 0.0001
Chronic steroid use (N, %)	5,335 (9%)	1,244 (17%)	4,091 (8%)	< 0.0001
Disseminated cancer (N, %)	2,704 (5%)	643 (9%)	2,061 (4%)	< 0.0001
Weight loss > 10% (N, %)	2,918 (5%)	802 (11%)	2,116 (4%)	< 0.0001
Bleeding disorder (N, %)	8,079 (14%)	1,929 (26%)	6,150 (12%)	< 0.0001
Active wound infection (N, %)	4,017 (7%)	997 (13%)	3,020 (6%)	< 0.0001
Preoperative sepsis (N, %)	29,541 (50%)	5,769 (77%)	23,772 (50%)	< 0.0001

N, number; AA, African American; CHF, congestive heart failure; COPD, chronic obstructive pulmonary disease; ESRD, end-stage renal disease.

TABLE 2. 30-Day Mortality Risk Calculator

Patient Factor	Patient Information	
Age	years	
Nonindependent functional status*	☐ Yes	
	□ No	
Preoperative sepsis**	☐ Yes	
	□ No	
BUN	mg/dL	
Albumin	g/dL	

Risk of 30-day mortality = 0.034*age + 0.8*nonindependent status + 0.88*sepsis + 1.1 (if BUN \geq 29) or 0.57 (if BUN \geq 18 and <29) + 1.16 (if albumin < 2.7) or 0.61 (if albumin \geq 2.7 and <3.4).

BUN, blood urea nitrogen.

albumin level (p < 0.001), a higher bilirubin level (p < 0.0001), a higher alkaline phosphatase level (p < 0.001), a higher partial thromboplastin time (PTT; p < 0.001), and a higher international normalized ratio (p < 0.0001) were significantly more likely to experience 30-day mortality. The patient characteristics and laboratory tests that were found to be most strongly associated with 30-day mortality in multivariate logistic regression were age, nonindependent functional status, the presence of preoperative sepsis, blood urea nitrogen level, and serum albumin level. Based on the strength of association of these variables with 30-day patient mortality, a mortality risk equation was generated as follows (Table 2):

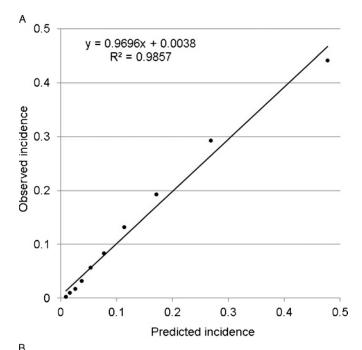
Risk of 30-day mortality = 0.034*age + 0.8*nonindependent status + 0.88*sepsis + 1.1 (if BUN ≥ 29) or 0.57 (if BUN ≥ 18 and <29) + 1.16 (if albumin < 2.7) or 0.61 (if albumin ≥ 2.7 and <3.4).

The discrimination of this model was good in both the derivation c statistic = 0.83 (95% confidence interval [CI], 0.82–0.83) and testing c statistic = 0.83 (95% CI, 0.82-0.84) data sets. In the validation data set, there was a strong linear association between predicted risk of 30-day mortality and actual incidence of 30-day mortality across risk deciles (Fig. 1A), with $R^2 = 0.99$. The regression line of predicted versus observed incidence had a slope of 0.97 and y-intercept of 0.004, indicating excellent model calibration. The same was true in the testing data (Fig. 1B), with $R^2 = 0.98$, slope = 0.96, and y-intercept = 0.005. The Pearson correlation between the model parameters in the derivation and testing data sets was r = 0.97, indicating good reproducibility of the risk calculator. As shown in Figure 2, the observed 30-day mortality in the testing group ranged from 2% (95% CI, 1%-2%) in the lowest decile to 31% (95% CI, 29%–33%) in the highest decile. A score ≤ 14.44 corresponded to decile one, >14.44-17.66 to decile two, >17.66-19.72 to decile three, >19.72-21.56 to decile four, >21.56-23.32 to decile five, >23.32-25.04 to decile six, >25.04-26.79 to decile seven, >26.79-28.55 to decile eight, >28.55-30.49 to decile nine, and >30.49 to decile ten. There was good discrimination between each risk decile with a narrow confidence interval for each decile, indicating high precision of the mortality risk calculator.

DISCUSSION

Several models have been proposed to help predict morbidity and mortality for particular cohorts of surgical patients.

For example, the American Society of Anesthesiologists developed the American Society of Anesthesiologists classification as a way to predict perioperative mortality for patients undergoing all types of surgery. The Goldman Index and the Revised Cardiac Index have been proposed to help predict the risk of cardiac complications and cardiac mortality in patients undergoing noncardiac surgery. The Acute Physiology and Chronic Health Evaluation (APACHE) II score was



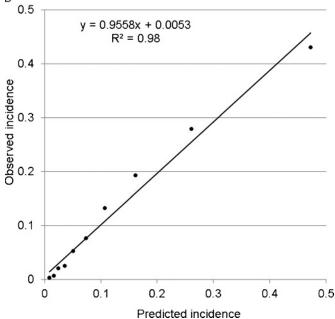


Figure 1. Calibration graph. With perfect calibration, the points should fall along a line with slope = 1 and y-intercept = 0. Each point represents the observed versus mean predicted probability of 30-day mortality within each risk score decile for (A) derivation sample and (B) testing sample.

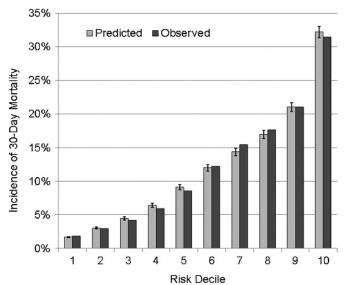


Figure 2. Predicted and observed incidence of 30-day mortality, stratified by risk decile, in the testing sample.

originally proposed for risk stratification of patients admitted to the intensive care unit but has since been adopted for the use in trauma surgery and other emergent general surgery cases. ¹⁷ Finally, an ACS-NSQIP surgical risk calculator was proposed in 2013 based on 21 patient variables found within the NSQIP database. ¹⁸ Although commonly used, all of these models have intrinsic limitations, either due to the use of subjective variables or the lack of generalizability to an EGS patient. ^{1,2,17,19–24}

The Affordable Care Act calls for transitioning from a volume-based physician reimbursement model to a quality-based model. Within this new model, outcomes are measured based on the ratio of the quality of care delivered to the cost required to deliver such care (value = quality/cost). Following suit to this rapidly evolving healthcare marketplace came the development of the ACS-NSQIP database. ACS-NSQIP is the first nationwide, validated, 30-day outcomes-centric database being used by academic, community, and academicaffiliated hospitals to measure quality of surgical care. The ACS-NSQIP includes surgical outcomes from over 600 hospitals and is believed to be representative of surgical practices throughout the United States.

The ACS-NSQIP published a universal surgical risk calculator in 2013. 18 This risk calculator uses 21 preoperative variables, both objective and subjective, to generate patient-specific estimates of 30-day perioperative morbidity and mortality. 18,30 Although it may seem peculiar that we used the ACS-NSQIP database to generate an EGS mortality risk calculator when there is already a universal ACS-NSQIP surgical risk calculator, a recent study by Bohnen et al.² demonstrates that perioperative morbidity and mortality for patients undergoing EGS may not be accurately predicted by the ACS-NSQIP surgical risk calculator. Furthermore, testing of the universal ACS-NSQIP surgical risk calculator during its derivation revealed that, when compared with procedure-specific ACS-NSOIP surgical risk calculators, the universal surgical risk calculator overestimates the risk of a morbidity and mortality event when an event is not observed and underestimates the risk of a

morbidity and mortality event when an event is observed. ¹⁸ Furthermore, the ACS-NSQIP surgical risk calculator allows the surgeon to adjust the morbidity and mortality risk of a patient up to two standard deviations higher than that generated by the risk calculator. ^{18,30} Finally, there are no studies that have validated the universal ACS-NSQIP surgical risk calculator. These concerns for the applicability of the ACS-NSQIP surgical risk calculator, coupled with the additional variability introduced by the surgeon adjustment factor, exposes the need for patient- and/or procedure-specific risk calculators such as the one we propose for patients undergoing EGS.

Because of the inherent emergent nature of EGS procedures, preoperative optimization is limited or nonexistent. This strengthens the clinical utility of a risk calculator in this patient population as the values used in the equation will remain relatively unchanged. Furthermore, since our risk calculator uses five preoperative variables currently included in routine evaluation, calculation of a patient's mortality risk can be easily done quickly and efficiently at the bedside without delay of surgical intervention. It is our intention that the use of this risk calculator will help to appropriately counsel and establish goals of care for the EGS patient and successfully allocate limited medical resources. For example, if a patient has a high risk of postoperative mortality and this is presented to them during preoperative discussion, they may choose to forgo surgical intervention. On the contrary, if they choose to proceed with surgical intervention, the surgical team can prepare the operating room team and intensive care unit with the same risk calculator information so that all available resources can be used to help minimize the risk of postoperative mortality.

In our study, we attempted to create a risk stratification model for 30-day mortality specifically for patients undergoing EGS due to the aforementioned lack of applicability of current risk stratification models to this patient population. 1,20-23 We chose to use only objective preoperative variables to strengthen the clinical utility of our model. With this goal in mind, our 30-day mortality risk score for patients undergoing EGS is based on five preoperative variables including age, nonindependent functional status, the presence or absence of preoperative sepsis, preoperative BUN level, and preoperative serum albumin level. Age and nonindependent functional status are components of the initial patient history taken during the preoperative examination. The presence of sepsis, as well as BUN and albumin, are almost ubiquitous components of the preoperative evaluation, whether in an academic or a rural setting. Our risk calculator, therefore, uses five readily available preoperative variables without the need for additional patient work-up or laboratory results that would otherwise delay surgical intervention. In other words, the use of variables inherent to a preoperative examination enhances the accuracy of this risk calculator due to limited circumstances in which these variables would be missing.

We recognize that there are limitations to our study. First and foremost, the applicability of our 30-day mortality risk calculator relies on external validation and comparisons to currently available mortality risk calculators. The most commonly used risk calculator for patients undergoing EGS is the APACHE II score. 17 Several of the variables used to calculate the APACHE II score are not available within the ACS-NSQIP database. Currently, the authors are working on a follow-up

study comparing the clinical accuracy of this 30-day mortality risk calculator in a cohort of EGS patients from a single institution with the APACHE II score. Second, our risk calculator depends on multiplication of factors based on the presence or absence of preoperative variables. Although this may limit the ability to use this calculator at the bedside, the use of a spreadsheet or other electronic tool such as a phone application for input of patient values may circumvent this challenge. Third, our risk calculator is limited to short-term mortality outcomes. Further studies are needed to determine intermediate (i.e., 90 days postoperatively) and long-term mortality outcomes after EGS. Finally, some may argue that our risk calculator is not applicable to non-NSQIP-participating hospitals. However, the ACS-NSQIP has expanded significantly since its inception, from 121 participating hospitals in 2005 to now 603 participating hospitals that include a wide variety of hospital settings.

CONCLUSION

Patients undergoing EGS constitute a unique patient population with limited applicability of previously established perioperative risk calculators. Our study found that risk stratification of this heterogeneous patient population can be performed with five readily available preoperative patient variables with relative ease. The clinical accuracy of this risk calculator requires additional prospective studies.

AUTHORSHIP

I.N.H., P.J.M., M.E.S., K.V., S.A., and B.S. conceived the study. I.N.H., R.L.A., and B.S. analyzed the data. I.N.H., P.J.M., M.E.S., R.L.A., K.V., and S.A. interpreted the data. I.N.H., P.J.M., and R.L.A. wrote the manuscript. I.N.H., M.E.S., R.L.A., K.V., S.A., and B.S. reviewed the final manuscript.

DISCLOSURE

American College of Surgeons National Surgical Quality Improvement Program and the hospitals participating in the ACS NSQIP are the source of the data used herein; they have not verified and are not responsible for the statistical validity of the data analysis or the conclusions derived by the authors.

REFERENCES

- Oliver CM, Walker E, Ginnaris S, Grocott MP, Moonesinghe SR. Risk assessment tools validated for patients undergoing emergency laparotomy: a systematic review. Br J Anaesth. 2015;115(6):849–860.
- Bohnen JD, Ramly EP, Sangji NF, de Moya M, Yeh DD, Lee J, Velmahos GC, Chang DC, Kaafarani HM. Perioperative risk factors impact outcomes in emergency versus nonemergency surgery differently: time to separate our national risk-adjustment models? *J Trauma Acute Care Surg*. 2016;81(1): 122–130.
- Gale SC, Shafi S, Dombrovskiy VY, Arumugam D, Crystal JS. The public health burden of emergency general surgery in the United States: a 10-year analysis of the nationwide inpatient sample—2001 to 2010. *J Trauma Acute* Care Surg. 2014;77(2):202–208.
- Scott JW, Olufajo OA, Brat GA, Rose JA, Zogg CK, Haider AH, Salim A, Havens JM. Use of national burden to define operative emergency general surgery. *JAMA Surg*. 2016;151(6):e160480.
- Havens JM, Peetz AB, Do WS, Cooper Z, Kelly E, Askari R, Reznor G, Salim A. The excess morbidity and mortality of emergency general surgery. J Trauma Acute Care Surg. 2015;78(2):306–311.
- Uno H, Cai T, Pencina MJ, D'Agostino RB, Wei LJ. On the C-statistics for evaluating overall adequacy of risk prediction procedures with censored survival data. Stat Med. 2011;30(10):1105–1117.
- Kim JH, Kim HS, Seo WY, Nam CM, Kim KY, Jeung HC, Lai JF, Chung HC, Noh SH, Rha SY. External validation of nomogram for the prediction

- of recurrence after curative resection in early gastric cancer. *Ann Oncol.* 2011;23(2):361–367.
- Ballian N, Rajamanickam V, Harms BA, Foley EF, Heise CP, Greenberg CC, Kennedy GD. Predictors of mortality after emergent surgery for acute colonic diverticulitis: analysis of national surgical quality improvement project data. J Trauma Acute Care Surg. 2012;74(2):611–616.
- Steyerberg EW, Vickers AJ, Cook NR, Gerds T, Gonen M, Obuchowski N, Pencina MJ, Kattan MW. Assessing the performance of prediction models: a framework for traditional and novel measures. *Epidemiology*. 2010;21(1): 128–138.
- ACS NSQIP Participant Use Data File. https://www.facs.org/qualityprograms/acs-nsqip/program-specifics/participant-use. Accessed April 21, 2016.
- Wolters U, Wolf T, Stützer H, Schröder T. ASA classification and perioperative variables as predictors of postoperative outcome. *Br J Anaesth*. 1996; 77(2):217–222.
- Saklad M. Grading of patients for surgical procedures. Anesthesiology. 1941; 2:281–284.
- Daabiss M. American Society of Anaesthesiologists Physical Status Classification. *Indian J Anaesth*. 2011;55(2):111–115.
- Glance LG, Lustik SJ, Hannan EL, Osler TM, Mukamel DB, Qian F, Dick AW. The surgical mortality probability model: derivation and validation of a simple risk prediction rule for noncardiac surgery. *Ann Surg*. 2012;255(4):696–702.
- Goldman L, Caldera DL, Nussbaum SR, Southwick FS, Krogstad D, Murray B, Burke DS, O'Malley TA, Goroll AH, Caplan CH, et al. Multifactorial index of cardiac risk in noncardiac surgical procedures. N Engl J Med. 1977;297(16):845–850.
- Lee TH, Marcantonio ER, Mangione CM, Thomas EJ, Polanczyk CA, Cook EF, Sugarbaker DJ, Donaldson MC, Poss R, Ho KK, et al. Derivation and prospective validation of a simple index for prediction of cardiac risk of major noncardiac surgery. *Circulation*. 1999;100(10):1043–1049.
- Goffi L, Saba V, Ghiselli R, Mattei A, Carle F. Preoperative Apache II and ASA scores in patients having major general surgical operations: prognostic value and potential clinical applications. *Eur J Surg.* 1999;165:730–735.
- Bilimoria KY, Liu Y, Paruch JL, Zhou L, Kmiecik TE, Ko CY, Cohen ME. Development and evaluation of the universal ACS NSQIP surgical risk calculator: a decision aid and informed consent tool for patients and surgeons. *J Am Coll Surg.* 2013;217(5):833–842.
- Riley R, Holman C, Fletcher D. Inter-rater reliability of the ASA physical status classification in a sample of anaesthetists in Western Australia. *Anaesth Intensive Care*. 2014;42(5):614

 –618.
- Shah N, Hamilton M. Clinical review: can we predict which patients are at risk of complications following surgery? Crit Care. 2013;17(3):226.
- Reilly BM, Evans AT. Translating clinical research into clinical practice: impact of using prediction rules to make decisions. *Ann Intern Med.* 2006; 144(3):201–209.
- Toll DB, Janssen KJ, Vergouwe Y, Moons KG. Validation, updating and impact of clinical prediction rules: a review. *J Clin Epidemiol*. 2008;61(11): 1085–1094.
- Moonesinghe SR, Mythen MG, Das P, Rowan KM, Grocott MP. Risk stratification tools for predicting morbidity and mortality in adult patients undergoing major surgery: qualitative systematic review. *Anesthesiology*. 2013; 119(4):959–981.
- Wyatt JC, Altman DG. Prognostic models: clinically useful or quickly forgotten. BMJ. 1995;311:1539–1541.
- Blumenthal D, Abrams M, Nuzum R. The Affordable Care Act at 5 years. N Engl J Med. 2015;372(25):2451–2458.
- 26. Porter ME. What is value in health care? N Engl J Med. 2010;363(26):2477–2481.
- Vogenberg FR. Predictive and prognostic models: implications for healthcare decision-making in modern recession. Am Health Drug Benefits. 2009;2(6): 218–222.
- ACS National Surgical Quality Improvement Program. https://www.facs.org/ quality-programs/acs-nsqip. Accessed April 13, 2016.
- Cevasco M, Ashley SW. Quality measurement and improvement in general surgery. Perm J. 2011;15(4):48–53.
- ACS NSQIP Surgical Risk Calculator. http://riskcalculator.facs.org/ RiskCalculator/PatientInfo.jsp. Accessed October 22, 2016.